

CS 147:
Computer Systems Performance Analysis
One-Factor Experiments

2015-06-15 CS147

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Computer Systems Performance Analysis
One-Factor Experiments

Overview

Introduction

The Model

Finding Effects

Calculating Errors

ANOVA

Allocation

Analysis

Verifying Assumptions

Unequal Sample Sizes

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└ Overview

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Verifying Assumptions

Unequal Sample Sizes

Characteristics of One-Factor Experiments

- ▶ Useful if there's only one important categorical factor with more than two interesting alternatives
 - ▶ Methods reduce to 2^1 factorial designs if only two choices
- ▶ If single variable isn't categorical, should use regression instead
- ▶ Method allows multiple replications

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└ Characteristics of One-Factor Experiments

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Comparing Truly Comparable Options

- ▶ Evaluating single workload on multiple machines
- ▶ Trying different options for single component
- ▶ Applying single suite of programs to different compilers

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└ Comparing Truly Comparable Options

Comparing Truly Comparable Options

- Evaluating single workload on multiple machines
- Trying different options for single component
- Applying single suite of programs to different compilers

When to Avoid It

- ▶ Incomparable “factors”
 - ▶ E.g., measuring vastly different workloads on single system
- ▶ Numerical factors
 - ▶ Won't predict any untested levels
 - ▶ Regression usually better choice
- ▶ Related entries across level
 - ▶ Use two-factor design instead

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An Example One-Factor Experiment

- ▶ Choosing authentication server for single-sized messages
- ▶ Four different servers are available
- ▶ Performance measured by response time
 - ▶ Lower is better

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An Example One-Factor Experiment

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The One-Factor Model

- ▶ $y_{ij} = \mu + \alpha_j + e_{ij}$
- ▶ y_{ij} is i^{th} response with factor set at level j
- ▶ μ is mean response
- ▶ α_j is effect of alternative j

$$\sum \alpha_j = 0$$

- ▶ e_{ij} is error term

$$\sum e_{ij} = 0$$

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└ The One-Factor Model

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$$\sum e_{ij} = 0$$

One-Factor Experiments With Replications

- ▶ Initially, assume r replications at each alternative of factor
- ▶ Assuming a alternatives, we have a total of ar observations
- ▶ Model is thus

$$\sum_{i=1}^r \sum_{j=1}^a y_{ij} = ar\mu + r \sum_{j=1}^a \alpha_j + \sum_{i=1}^r \sum_{j=1}^a e_{ij}$$

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Sample Data for Our Example

- ▶ Four alternatives, with four replications each (measured in seconds)

A	B	C	D
0.96	0.75	1.01	0.93
1.05	1.22	0.89	1.02
0.82	1.13	0.94	1.06
0.94	0.98	1.38	1.21

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Computing Effects

- ▶ Need to figure out μ and α_j
- ▶ We have various y_{ij} 's
- ▶ Errors should add to zero:

$$\sum_{i=1}^r \sum_{j=1}^a e_{ij} = 0$$

- ▶ Similarly, effects should add to zero:

$$\sum_{j=1}^a \alpha_j = 0$$

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└ Computing Effects

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Calculating μ

- ▶ By definition, sum of errors and sum of effects are both zero:

$$\sum_{i=1}^r \sum_{j=1}^a y_{ij} = ar\mu + 0 + 0$$

- ▶ And thus, μ is equal to **grand mean** of all responses

$$\mu = \frac{1}{ar} \sum_{i=1}^r \sum_{j=1}^a y_{ij} = \bar{y}_{..}$$

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└ Finding Effects

└ Calculating μ Calculating μ

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Calculating μ for Our Example

Thus,

$$\begin{aligned}\mu &= \frac{1}{4 \times 4} \sum_{i=1}^4 \sum_{j=1}^4 y_{ij} \\ &= \frac{1}{16} \times 16.29 \\ &= 1.018\end{aligned}$$

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└ Finding Effects

└ Calculating μ for Our ExampleCalculating μ for Our Example

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Calculating α_j

- ▶ α_j is vector of responses
 - ▶ One for each alternative of the factor
- ▶ To find vector, find **column means**

$$\bar{y}_{.j} = \frac{1}{r} \sum_{i=1}^r y_{ij}$$

- ▶ Separate mean for each j
- ▶ Can calculate directly from observations

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└ Finding Effects

└ Calculating α_j Calculating α_j

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Calculating Column Mean

- ▶ We know that y_{ij} is defined to be

$$y_{ij} = \mu + \alpha_j + \mathbf{e}_{ij}$$

- ▶ So,

$$\begin{aligned} \bar{y}_{.j} &= \frac{1}{r} \sum_{i=1}^r (\mu + \alpha_j + \mathbf{e}_{ij}) \\ &= \frac{1}{r} \left(r\mu + r\alpha_j + \sum_{i=1}^r \mathbf{e}_{ij} \right) \end{aligned}$$

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└ Finding Effects

└ Calculating Column Mean

Calculating Column Mean

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Calculating Parameters

- ▶ Sum of errors for any given row is zero, so

$$\begin{aligned}\bar{y}_{.j} &= \frac{1}{r}(r\mu + r\alpha_j + 0) \\ &= \mu + \alpha_j\end{aligned}$$

- ▶ So we can solve for α_j :

$$\alpha_j = \bar{y}_{.j} - \mu = \bar{y}_{.j} - \bar{y}_{..}$$

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Parameters for Our Example

Server	A	B	C	D
Col. Mean	.9425	1.02	1.055	1.055

Subtract μ from column means to get parameters:

Parameters	-.076	.002	.037	.037
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 └─ Parameters for Our Example

Parameters for Our Example

Server	A	B	C	D
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Estimating Experimental Errors

- ▶ Estimated response is $\hat{y}_{ij} = \mu + \alpha_{ij}$
- ▶ But we measured actual responses
 - ▶ Multiple responses per alternative
- ▶ So we can estimate amount of error in estimated response
- ▶ Use methods similar to those used in other types of experiment designs

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└ Calculating Errors

└ Estimating Experimental Errors

Estimating Experimental Errors

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Sum of Squared Errors

- ▶ SSE estimates variance of the errors:

$$\text{SSE} = \sum_{i=1}^r \sum_{j=1}^a e_{ij}^2$$

- ▶ We can calculate SSE directly from model and observations
- ▶ Also can find indirectly from its relationship to other error terms

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SSE for Our Example

Calculated directly:

$$\begin{aligned}
 SSE &= (.96 - (1.018 - .076))^2 \\
 &\quad + (1.05 - (1.018 - .076))^2 + \dots \\
 &\quad + (.75 - (1.018 + .002))^2 \\
 &\quad + (1.22 - (1.018 + .002))^2 + \dots \\
 &\quad + (.93 - (1.018 + .037))^2 \\
 &= .3425
 \end{aligned}$$

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 - └ Calculating Errors
 - └ SSE for Our Example

SSE for Our Example

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 SSE &= (.96 - (1.018 - .076))^2 \\
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 &\quad + (.93 - (1.018 + .037))^2 \\
 &= .3425
 \end{aligned}$$

Allocating Variation

- ▶ To allocate variation for model, start by squaring both sides of model equation

$$y_{ij}^2 = \mu^2 + \alpha_j^2 + e_{ij}^2 + 2\mu\alpha_j + 2\mu e_{ij} + 2\alpha_j e_{ij}$$

$$\sum_{i,j} y_{ij}^2 = \sum_{i,j} \mu^2 + \sum_{i,j} \alpha_j^2 + \sum_{i,j} e_{ij}^2$$

+ cross-products

- ▶ Cross-product terms add up to zero

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Variation In Sum of Squares Terms

$$SSY = SS0 + SSA + SSE$$

$$SSY = \sum_{i,j} y_{ij}^2$$

$$SS0 = \sum_{i=1}^r \sum_{j=1}^a \mu^2 = ar\mu^2$$

$$SSA = \sum_{i=1}^r \sum_{j=1}^a \alpha_j^2 = r \sum_{j=1}^a \alpha_j^2$$

Gives another way to calculate SSE

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└ Variation In Sum of Squares Terms

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Gives another way to calculate SSE

Sum of Squares Terms for Our Example

- ▶ $SSY = 16.9615$
- ▶ $SS0 = 16.58256$
- ▶ $SSA = .03377$
- ▶ So SSE must equal $16.9615 - 16.58256 - .03377$
 - ▶ $= 0.3425$
 - ▶ Matches our earlier SSE calculation

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└ Sum of Squares Terms for Our Example

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Assigning Variation

- ▶ SST is total variation
- ▶ $SST = SSY - SS0 = SSA + SSE$
- ▶ Part of total variation comes from model
- ▶ Part comes from experimental errors
- ▶ A good model explains a lot of variation

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└ Assigning Variation

Assigning Variation

- SST is total variation
- $SST = SSY - SS0 = SSA + SSE$
- Part of total variation comes from model
- Part comes from experimental errors
- A good model explains a lot of variation

Assigning Variation in Our Example

- ▶ $SST = SSY - SS0 = 0.376244$
- ▶ $SSA = .03377$
- ▶ $SSE = .3425$
- ▶ Percentage of variation explained by server choice:

$$= 100 \times \frac{.03377}{.3762} = 8.97\%$$

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 └ Assigning Variation in Our Example

Assigning Variation in Our Example

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Analysis of Variance

- ▶ Percentage of variation explained can be large or small
- ▶ Regardless of size, may or may not be statistically significant
- ▶ To determine significance, use ANOVA procedure
 - ▶ Assumes normally distributed errors

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└ Analysis of Variance

Analysis of Variance

- Percentage of variation explained can be large or small
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Running ANOVA

- ▶ Easiest to set up tabular method
- ▶ Like method used in regression models
 - ▶ Only slight differences
- ▶ Basically, determine ratio of Mean Squared of A (parameters) to Mean Squared Errors
- ▶ Then check against F-table value for number of degrees of freedom

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Running ANOVA

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ANOVA Table for One-Factor Experiments

Component	Sum of Squares	% of Variation	Degrees of Freedom	Mean Square	F-Computed	F-Table
y	$SSY = \sum y_{ij}^2$		N			
$\bar{y}_{..}$	$SS0 = N\mu^2$		1			
$y - \bar{y}_{..}$	$SST = SSY - SS0$	100	$N - 1$			
A	$SSA = r \sum \alpha_j^2$	$\frac{SSA}{SST}$	$a - 1$	$MSA = \frac{SSA}{a-1}$	$\frac{MSA}{MSE}$	F[$1 - \alpha$; $a - 1$, $N - a$]
e	$SSE = SST - SSA$	$\frac{SSE}{SST}$	$N - a$	$MSE = \frac{SSE}{N-a}$		

$$N = ar \quad s_e = \sqrt{MSE}$$

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└ ANOVA

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A	$SSA = r \sum \alpha_j^2$	$\frac{SSA}{SST}$	$a - 1$	$MSA = \frac{SSA}{a-1}$	$\frac{MSA}{MSE}$	F[$1 - \alpha$; $a - 1$, $N - a$]
e	$SSE = SST - SSA$	$\frac{SSE}{SST}$	$N - a$	$MSE = \frac{SSE}{N-a}$		

ANOVA Procedure for Our Example

Component	Sum of Squares	% of Variation	Degrees of Freedom	Mean Square	F-Computed	F-Table
y	16.96		16			
$\bar{y}..$	16.58		1			
$y - \bar{y}..$	0.376	100	15			
A	.034	9.0	3	.011	0.394	2.61
e	.342	91.0	12	.028		

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└ ANOVA Procedure for Our Example

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Interpretation of Sample ANOVA

- ▶ Done at 90% level
- ▶ F-computed is .394
- ▶ Table entry at 90% level with $n = 3$ and $m = 12$ is 2.61
- ▶ Thus, servers are not significantly different

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Interpretation of Sample ANOVA

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- F-computed is .394
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One-Factor Experiment Assumptions

- ▶ Analysis of one-factor experiments makes the usual assumptions:
 - ▶ Effects of factors are additive
 - ▶ Errors are additive
 - ▶ Errors are independent of factor alternatives
 - ▶ Errors are normally distributed
 - ▶ Errors have same variance at all alternatives
- ▶ How do we tell if these are correct?

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└ One-Factor Experiment Assumptions

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Visual Diagnostic Tests

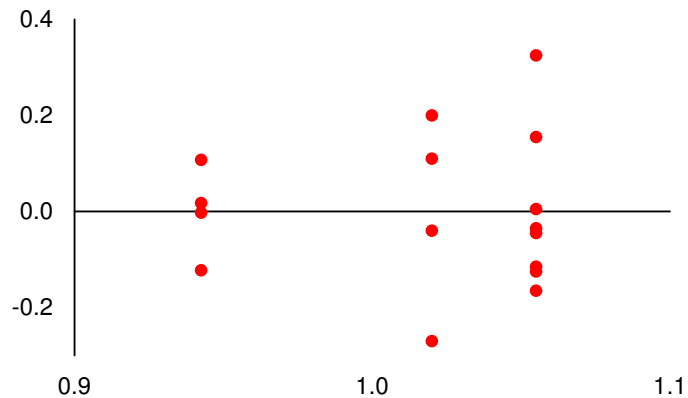
- ▶ Similar to those done before
 - ▶ Residuals vs. predicted response
 - ▶ Normal quantile-quantile plot
 - ▶ Residuals vs. experiment number

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└ Visual Diagnostic Tests

Visual Diagnostic Tests

- Similar to those done before
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 - Normal quantile-quantile plot
 - Residuals vs. experiment number

Residuals vs. Predicted for Example



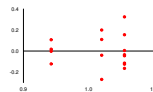
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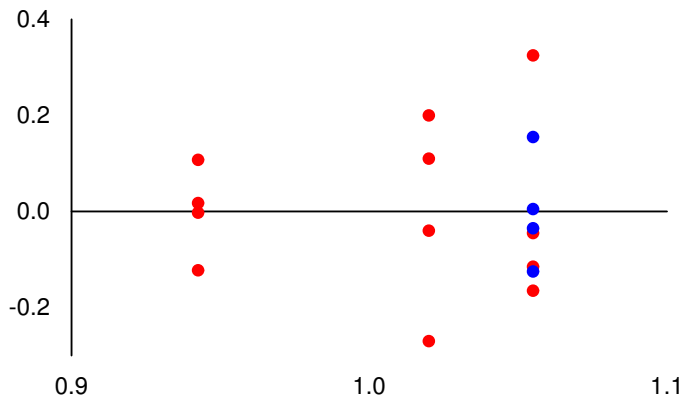
└ Verifying Assumptions

└ Residuals vs. Predicted for Example

Residuals vs. Predicted for Example



Residuals vs. Predicted, Slightly Revised



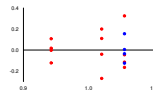
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└ Verifying Assumptions

└ Residuals vs. Predicted, Slightly Revised

Residuals vs. Predicted, Slightly Revised



In the alternate rendering, the predictions for server D are shown in blue so they can be distinguished from server C.

What Does The Plot Tell Us?

- ▶ Analysis assumed size of errors was unrelated to factor alternatives
- ▶ Plot tells us something entirely different
 - ▶ Very different spread of residuals for different factors
- ▶ Thus, one-factor analysis is not appropriate for this data
 - ▶ Compare individual alternatives instead
 - ▶ Use pairwise confidence intervals

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└ Verifying Assumptions

└ What Does The Plot Tell Us?

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Could We Have Figured This Out Sooner?

- ▶ Yes!
- ▶ Look at original data
- ▶ Look at calculated parameters
- ▶ Model says C & D are identical
- ▶ Even cursory examination of data suggests otherwise

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└ Verifying Assumptions

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Looking Back at the Data

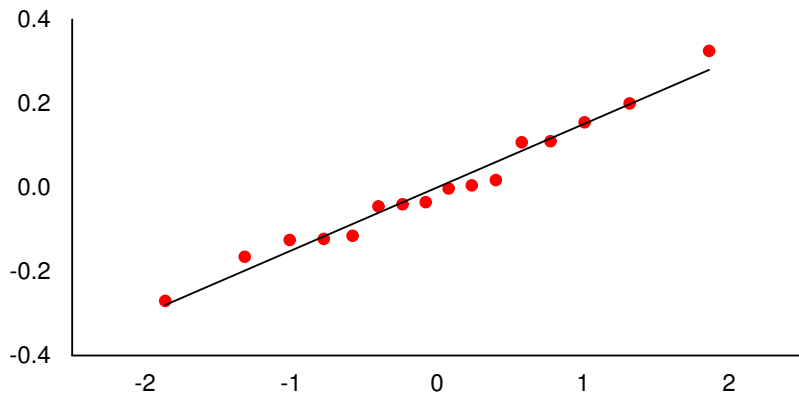
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Looking Back at the Data

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Quantile-Quantile Plot for Example



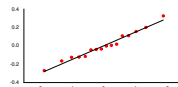
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└ Quantile-Quantile Plot for Example

Quantile-Quantile Plot for Example



What Does This Plot Tell Us?

- ▶ Overall, errors are normally distributed
- ▶ If we only did quantile-quantile plot, we'd think everything was fine
- ▶ The lesson: **test ALL assumptions, not just one or two**

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└ Verifying Assumptions

└ What Does This Plot Tell Us?

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- The lesson: **test ALL assumptions, not just one or two**

One-Factor Confidence Intervals

- ▶ Estimated parameters are random variables
 - ▶ Thus, can compute confidence intervals
- ▶ Basic method is same as for confidence intervals on $2^k r$ design effects
- ▶ Find standard deviation of parameters
 - ▶ Use that to calculate confidence intervals
 - ▶ Possible typo in book, p. 336, example 20.6, in formula for calculating α_j
 - ▶ Also might be typo on p. 335: degrees of freedom is $a(r - 1)$, not $r(a - 1)$

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└ Verifying Assumptions

└ One-Factor Confidence Intervals

One-Factor Confidence Intervals

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 - Use that to calculate confidence intervals
 - Possible typo in book, p. 336, example 20.6, in formula for calculating α_j
 - Also might be typo on p. 335: degrees of freedom is $a(r - 1)$, not $r(a - 1)$

Confidence Intervals For Example Parameters

- ▶ $s_e = .158$
- ▶ Standard deviation of $\mu = .040$
- ▶ Standard deviation of $\alpha_j = .069$
- ▶ 95% confidence interval for $\mu = (.932, 1.10)$
- ▶ 95% CI for $\alpha_1 = (-.225, .074)$
- ▶ 95% CI for $\alpha_2 = (-.148, .151)$
- ▶ 95% CI for $\alpha_3 = (-.113, .186)$
- ▶ 95% CI for $\alpha_4 = (-.113, .186)$

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└ Verifying Assumptions

└ Confidence Intervals For Example Parameters

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Unequal Sample Sizes in One-Factor Experiments

- ▶ Don't really need identical replications for all alternatives
- ▶ Only slight extra difficulty
- ▶ See book example for full details

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└ Unequal Sample Sizes

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Changes To Handle Unequal Sample Sizes

- ▶ Model is the same
- ▶ Effects are weighted by number of replications for that alternative:

$$\sum_{j=1}^a r_j a_j = 0$$

- ▶ Slightly different formulas for degrees of freedom

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└ Changes To Handle Unequal Sample Sizes

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