

CS 147:
Computer Systems Performance Analysis
Two-Factor Designs

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Computer Systems Performance Analysis
Two-Factor Designs

Overview

Two-Factor Designs

No Replications

Adding Replications

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└ Overview

Overview

Two-Factor Designs
No Replications
Adding Replications

Two-Factor Design Without Replications

- ▶ Used when only two parameters, but multiple levels for each
- ▶ Test all combinations of levels of the two parameters
- ▶ One replication (observation) per combination
- ▶ For factors A and B with a and b levels, ab experiments required

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└ Two-Factor Designs
└ No Replications
└ Two-Factor Design Without Replications

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- Test all combinations of levels of the two parameters
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When to Use This Design?

- ▶ System has two important factors
- ▶ Factors are categorical
- ▶ More than two levels for at least one factor
- ▶ Examples:
 - ▶ Performance of different processors under different workloads
 - ▶ Characteristics of different compilers for different benchmarks
 - ▶ Performance of different Web browsers on different sites

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└ When to Use This Design?

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 - Performance of different processors under different workloads
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When to Avoid This Design?

- ▶ Systems with more than two important factors
 - ▶ Use general factorial design
- ▶ Non-categorical variables
 - ▶ Use regression
- ▶ Only two levels per factor
 - ▶ Use 2^2 designs

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└ When to Avoid This Design?

When to Avoid This Design?

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 - Use general factorial design
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 - Use regression
- Only two levels per factor
 - Use 2^2 designs

Model For This Design

- ▶ $y_{ij} = \mu + \alpha_j + \beta_i + e_{ij}$
- ▶ y_{ij} is observation
- ▶ μ is mean response
- ▶ α_j is effect of factor A at level j
- ▶ β_i is effect of factor B at level i
- ▶ e_{ij} is error term
- ▶ Sums of α_j 's and β_i 's are both zero

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└ Model For This Design

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Assumptions of the Model

- ▶ Factors are additive
- ▶ Errors are additive
- ▶ Typical assumptions about errors:
 - ▶ Distributed independently of factor levels
 - ▶ Normally distributed
- ▶ Remember to check these assumptions!

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└ Assumptions of the Model

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Computing Effects

- ▶ Need to figure out μ , α_j , and β_i
- ▶ Arrange observations in two-dimensional matrix
 - ▶ b rows, a columns
- ▶ Compute effects such that error has zero mean
 - ▶ Sum of error terms across all rows and columns is zero

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└ No Replications
└ Computing Effects

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Two-Factor Full Factorial Example

- ▶ Want to expand functionality of a file system to allow automatic compression
- ▶ Examine three choices:
 - ▶ Library substitution of file system calls
 - ▶ New VFS
 - ▶ Stackable layers
- ▶ Three different benchmarks
- ▶ Metric: response time

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└ Two-Factor Full Factorial Example

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Data for Example

	<u>Library</u>	<u>VFS</u>	<u>Layers</u>
Compile Benchmark	94.3	89.5	96.2
Email Benchmark	224.9	231.8	247.2
Web Server Benchmark	733.5	702.1	797.4

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Data for Example

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Compile Benchmark	94.3	89.5	96.2
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Computing μ

- ▶ Averaging the j^{th} column,

$$\bar{y}_{.j} = \mu + \alpha_j + \frac{1}{b} \sum_i \beta_i + \frac{1}{b} \sum_i e_{ij}$$

- ▶ By assumption, error terms add to zero
- ▶ Also, the β_j 's add to zero, so $\bar{y}_{.j} = \mu + \alpha_j$
- ▶ Averaging rows produces $\bar{y}_{i.} = \mu + \beta_i$
- ▶ Averaging everything produces $\bar{y}_{..} = \mu$

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 └ Computing μ

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Model Parameters

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└ Two-Factor Designs
└ No Replications
└ Model Parameters

Model Parameters

Using same techniques as for one-factor designs, parameters are:

- $\bar{y}_{..} = \mu$
- $\alpha_j = \bar{y}_{.j} - \bar{y}_{..}$
- $\beta_i = \bar{y}_{i.} - \bar{y}_{..}$

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- ▶ $\bar{y}_{..} = \mu$
- ▶ $\alpha_j = \bar{y}_{.j} - \bar{y}_{..}$
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Calculating Parameters for the Example

- ▶ μ = grand mean = 357.4
- ▶ $\alpha_j = (-6.5, -16.3, 22.8)$
- ▶ $\beta_i = (-264.1, -122.8, 386.9)$
- ▶ So, for example, the model predicts that the email benchmark using a special-purpose VFS will take $357.4 - 16.3 - 122.8 = 218.3$ seconds

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└ Calculating Parameters for the Example

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Estimating Experimental Errors

- ▶ Similar to estimation of errors in previous designs
- ▶ Take difference between model's predictions and observations
- ▶ Calculate Sum of Squared Errors
- ▶ Then allocate variation

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└ Two-Factor Designs
└ No Replications
└ Estimating Experimental Errors

Estimating Experimental Errors

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Allocating Variation

- ▶ Use same kind of procedure as on other models
- ▶ $SSY = SS0 + SSA + SSB + SSE$
- ▶ $SST = SSY - SS0$
- ▶ Can then divide total variation between SSA, SSB, and SSE

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└ Two-Factor Designs
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└ Allocating Variation

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Calculating SS0, SSA, SSB

- ▶ $SS0 = ab\mu^2$
- ▶ $SSA = b \sum_j \alpha_j^2$
- ▶ $SSB = a \sum_i \beta_i^2$
- ▶ Recall that a and b are numbers of levels for the factors

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└ No Replications
└ Calculating SS0, SSA, SSB

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- Recall that a and b are numbers of levels for the factors

Allocation of Variation for Example

- ▶ $SSE = 2512$
- ▶ $SSY = 1,858,390$
- ▶ $SS0 = 1,149,827$
- ▶ $SSA = 2489$
- ▶ $SSB = 703,561$
- ▶ $SST = 708,562$
- ▶ Percent variation due to A: 0.35%
- ▶ Percent variation due to B: 99.3%
- ▶ Percent variation due to errors: 0.35%

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Analysis of Variation

- ▶ Again, similar to previous models, with slight modifications
- ▶ As before, use an ANOVA procedure
 - ▶ Need extra row for second factor
 - ▶ Minor changes in degrees of freedom
- ▶ End steps are the same
 - ▶ Compare F-computed to F-table
 - ▶ Compare for each factor

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Analysis of Variation for Our Example

- ▶ $MSE = SSE / [(a - 1)(b - 1)] = 2512 / [(2)(2)] = 628$
- ▶ $MSA = SSA / (a - 1) = 2489 / 2 = 1244$
- ▶ $MSB = SSB / (b - 1) = 703,561 / 2 = 351,780$
- ▶ F-computed for A = $MSA / MSE = 1.98$
- ▶ F-computed for B = $MSB / MSE = 560$
- ▶ 95% F-table value for A & B is 6.94
- ▶ So A is not significant, but B is

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└ Analysis of Variation for Our Example

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Checking Our Results with Visual Tests

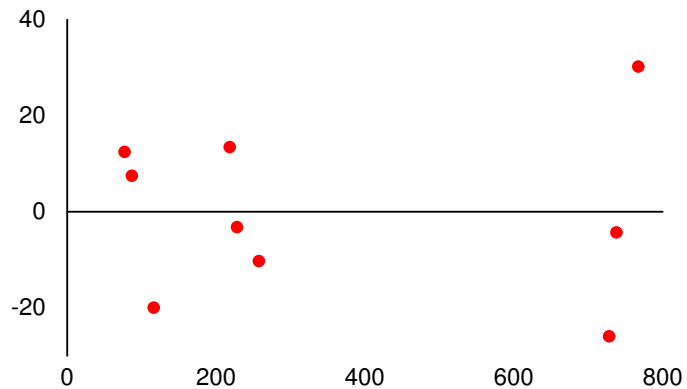
- ▶ As always, check if assumptions made in the analysis are correct
- ▶ Use residuals vs. predicted and quantile-quantile plots

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└ No Replications
└ Checking Our Results with Visual Tests

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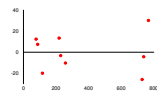
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Residuals vs. Predicted Response for Example



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└ Residuals vs. Predicted Response for Example

Residuals vs. Predicted Response for Example



What Does the Chart Reveal?

- ▶ Do we or don't we see a trend in errors?
- ▶ Clearly they're higher at highest level of the predictors
- ▶ But is that alone enough to call a trend?
 - ▶ Perhaps not, but we should take a close look at both factors to see if there's reason to look further
 - ▶ Maybe take results with a grain of salt

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└ Two-Factor Designs

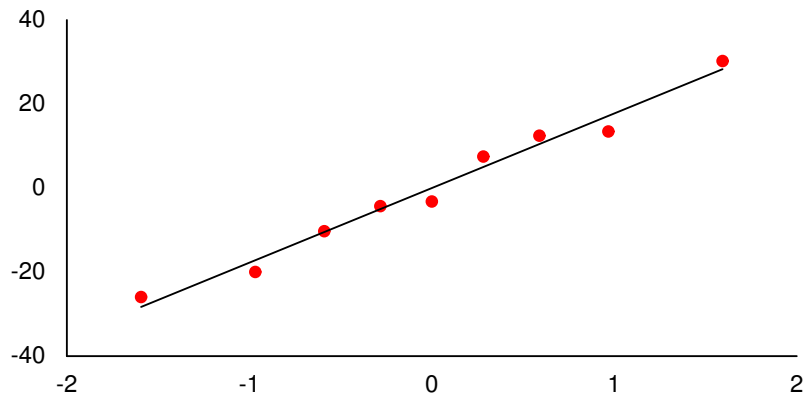
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Quantile-Quantile Plot for Example



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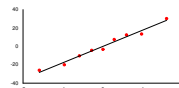
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└ Quantile-Quantile Plot for Example

Quantile-Quantile Plot for Example



Confidence Intervals for Effects

- ▶ Need to determine standard deviation for data as a whole
- ▶ Then can derive standard deviations for effects
 - ▶ Use different degrees of freedom for each
- ▶ Complete table in Jain, p. 351

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└ Confidence Intervals for Effects

Confidence Intervals for Effects

- Need to determine standard deviation for data as a whole
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- Complete table in Jain, p. 351

Standard Deviations for Example

- ▶ $s_e = 25$
- ▶ Standard deviation of μ :

$$s_\mu = s_e / \sqrt{ab} = 25 / \sqrt{3 \times 3} = 8.3$$

- ▶ Standard deviation of α_j :

$$s_{\alpha_j} = s_e \sqrt{(a-1)/ab} = 25 \sqrt{2/9} = 11.8$$

- ▶ Standard deviation of β_i :

$$s_{\beta_i} = s_e \sqrt{(b-1)/ab} = 25 \sqrt{2/9} = 11.8$$

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Calculating Confidence Intervals for Example

- ▶ Only file system alternatives shown here
- ▶ We'll use 95% level
- ▶ 4 degrees of freedom
- ▶ CI for library solution: $(-39, 26)$
- ▶ CI for VFS solution: $(-49, 16)$
- ▶ CI for layered solution: $(-10, 55)$
- ▶ So none of the solutions are significantly different from mean at 95% confidence

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└ Calculating Confidence Intervals for Example

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Looking a Little Closer

- ▶ Do zero CI's mean that none of the alternatives for adding functionality are different?
- ▶ Not necessarily
- ▶ Use contrasts to check (see Section 18.5 & p. 366)

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└ Looking a Little Closer

Looking a Little Closer

- Do zero CIs mean that none of the alternatives for adding functionality are different?
- Not necessarily
- Use contrasts to check (see Section 18.5 & p. 366)

Comparing Contrasts

- ▶ Is library approach significantly better than layers?
- ▶ Define a contrast: $u = \sum_{j=1}^a h_j \alpha_j$ where h_j 's are chosen so that $\sum_{j=1}^a h_j = 0$
- ▶ To compare library vs. layers, set $h = (1, 0, -1)$
- ▶ Contrast mean = $\sum_{j=1}^a h_j \bar{y}_{.j} = 350.9 - 380.267 = -29.367$
- ▶ Contrast variance = $s_e^2(\sum_{j=1}^a h_j^2)/b = 25 \times 2/3 = 16.667$, so contrast s.d. = 4.082
- ▶ Using $t_{[1-\alpha/2; (a-1)(b-1)]} = t_{[.975; 4]} = 2.776$, confidence interval is $-29.367 \mp 4.082 \times 2.776 = (-40.7, -18.0)$
- ▶ So library approach is better, at 95%

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 - └ Comparing Contrasts

Comparing Contrasts

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- So library approach is better, at 95%

Missing Observations

- ▶ Sometimes experiments go awry
- ▶ You don't want to discard an entire study away just because one observation got lost
- ▶ Solution:
 - ▶ Calculate row/column means and standard deviations based on actual observation count
 - ▶ Degrees of freedom in SS^* also must be adjusted
 - ▶ See book for example
- ▶ Alternatives exist but are controversial
- ▶ If lots of missing values in a column or row, throw it out entirely
 - ▶ Best is to have only 1–2 missing values

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Replicated Two-Factor Designs

- ▶ For r replications of each experiment, model becomes

$$y_{ijk} = \mu + \alpha_j + \beta_i + \gamma_{ij} + e_{ijk}$$

- ▶ γ_{ij} represents interaction between factor A at level j and B at level i
- ▶ As before, effect sums $\sum \alpha_j$ and $\sum \beta_i$ are zero
- ▶ Interactions are zero for both row and column sums:

$$\forall i \quad \sum_{j=1}^a \gamma_{ij} = 0 \quad \forall j \quad \sum_{i=1}^b \gamma_{ij} = 0$$

- ▶ Per-experiment errors add to zero:

$$\forall i, j \quad \sum_{k=1}^r e_{ijk} = 0$$

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└ Replicated Two-Factor Designs

Replicated Two-Factor Designs

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- Per-experiment errors add to zero:

$$\forall i, j \quad \sum_{k=1}^r e_{ijk} = 0$$

Calculating Effects

Same as usual:

- ▶ Calculate grand mean $\bar{y}_{...}$, row and column means $\bar{y}_{i..}$ and $\bar{y}_{.j}$, and per-experiment means \bar{y}_{ij} .
- ▶ $\mu = \bar{y}_{...}$
- ▶ $\alpha_j = \bar{y}_{.j} - \mu$
- ▶ $\beta_i = \bar{y}_{i..} - \mu$
- ▶ $\gamma_{ij} = \bar{y}_{ij} - \alpha_j - \beta_i - \mu$
- ▶ $e_{ijk} = y_{ijk} - \bar{y}_{ij}$.

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- └ Two-Factor Designs
 - └ Adding Replications
 - └ Calculating Effects

Calculating Effects

Same as usual:

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- $e_{ijk} = y_{ijk} - \bar{y}_{ij}$

Analysis of Variance

- ▶ Again, extension of earlier models
- ▶ See Table 22.5, p. 375, for formulas
- ▶ As usual, must do visual tests

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└ Analysis of Variance

Analysis of Variance

- Again, extension of earlier models
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Why Can We Find Interactions?

- ▶ Without replications, two-factor model didn't give interactions
- ▶ Why not?

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This slide has animations.

In unreplicated experiment, we could have assumed no experimental errors and attributed variation to interaction instead (but that wouldn't be wise).

Why Can We Find Interactions?

- ▶ Without replications, two-factor model didn't give interactions
- ▶ Why not?
- ▶ Insufficient data
- ▶ Variation from predictions was attributed to errors, not interaction
 - ▶ Interaction is confounded with errors
- ▶ Now, we have more info
 - ▶ For given A , B setting, errors are assumed to cause variation in r replicated experiments
 - ▶ Any remaining variation must therefore be interaction

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General Full Factorial Designs

- ▶ Straightforward extension of two-factor designs
- ▶ Average along axes to get effects
- ▶ Must consider all interactions (various axis combinations)
- ▶ Regression possible for quantitative effects
 - ▶ But should have more than three data points
- ▶ If no replications, errors confounded with highest-level interaction

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└ General Full Factorial Designs

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